



TITLE:

# A new prognostic index for overall survival in malignant pleural mesothelioma: the rPHS (regimen, PS, histology or stage) index.

AUTHOR(S):

Kataoka, Yuki; Yamamoto, Yosuke; Otsuki, Taiichiro;  
Shinomiya, Mariko; Terada, Takayuki; Fukuma, Shingo;  
Yamazaki, Shin; Hirabayashi, Masataka; Nakano, Takashi;  
Fukuhara, Shunichi

---

CITATION:

Kataoka, Yuki ...[et al]. A new prognostic index for overall survival in malignant pleural mesothelioma: the rPHS (regimen, PS, histology or stage) index.. Japanese journal of clinical oncology 2015, 45(6): 562-568

ISSUE DATE:

2015-06

URL:

<http://hdl.handle.net/2433/202002>

RIGHT:

This is a pre-copyedited, author-produced PDF of an article accepted for publication in 'Japanese journal of clinical oncology' following peer review. The version of record [Yuki Kataoka, Yosuke Yamamoto, Taiichiro Otsuki, Mariko Shinomiya, Takayuki Terada, Shingo Fukuma, Shin Yamazaki, Masataka Hirabayashi, Takashi Nakano, Shunichi Fukuhara. A new prognostic index for overall survival in malignant pleural mesothelioma: the rPHS (regimen, PS, histology or stage) index. Japanese Journal of Clinical Oncology (2015) 45 (6): 562-568. doi: 10.1093/jjco/hyv039] is available online at: <http://jjco.oxfordjournals.org/content/45/6/562>.; This is not the published version. Please cite only the published version.; この論文は出版社版ではありません。引用の際には出版社版をご確認ご利用ください。

1 A new prognostic index for overall survival in malignant pleural mesothelioma: the  
2 rPHS (regimen, PS, Histology, or Stage) index

3

4 Yuki Kataoka<sup>1,3</sup>, Yosuke Yamamoto<sup>1</sup>, Taiichiro Otsuki<sup>2</sup>, Mariko Shinomiya<sup>3</sup>, Takayuki  
5 Terada<sup>4</sup>, Shingo Fukuma<sup>1,5</sup>, Shin Yamazaki<sup>1</sup>, Masataka Hirabayashi<sup>3</sup>, Takashi Nakano<sup>4</sup>,  
6 Shunichi Fukuhara<sup>1,5,\*</sup>

7

8 1 Department of Healthcare Epidemiology, Graduate School of Medicine and Public  
9 Health, Kyoto University, Yoshida Konoe-cho, Sakyo-ku, Kyoto, 2 Cancer center,  
10 Hyogo College of Medicine, Mukogawa-cho, Nishinomiya, Hyogo, 3 Department of  
11 Respiratory Medicine, Hyogo Prefectural Amagasaki Hospital, Higashi-Daimotsu-Cho,  
12 Amagasaki, Hyogo 4 Division of Respiratory Medicine, Hyogo College of Medicine,  
13 Hyogo, Japan, Hyogo College of Medicine, Mukogawa-cho, Nishinomiya, Hyogo, and  
14 5 Center for Innovative Research in Clinical Evaluative Science (CiRCLE), Fukushima  
15 Medical University, Hikarigaoka, Fukushima

16

17 \*For reprints and all correspondence:

18 Shunichi Fukuhara, MD, DMSc, Department of Healthcare Epidemiology, Graduate

19 School of Medicine and Public Health, Kyoto University, Yoshida Konoe-cho, Sakyo-ku,

20 Kyoto 606-8501, Japan.

21 Tel: +81-75-753-4646 Fax: +81-75-753-4644

22 Email: [fukuhara.shunichi.6m@kyoto-u.ac.jp](mailto:fukuhara.shunichi.6m@kyoto-u.ac.jp)

23

24 Running head:

25 A prognostic index in malignant mesothelioma

26

27

## 28 ABSTRACT

### 29 Background

30 Existing prognostic indices (PI) for malignant pleural mesothelioma (MPM) do not  
31 incorporate the recent advances in oncology care. The purpose of this study was to  
32 provide a PI for overall survival (OS) in MPM patients treated with chemotherapy with  
33 pemetrexed (PEM) or best supportive care (BSC) in the recent clinical setting.

### 34 Methods

35 A retrospective cohort study was performed in two hospitals in Japan (2007 - 2013).  
36 The primary outcomes were OS. The Cox proportional hazards model was used for  
37 multivariable analyses to identify prognostic factors. A final model was chosen based on  
38 both clinical and statistical significance.

### 39 Results

40 A total of 283 patients (CTx: n=228, BSC: n=55) were enrolled in the study. On  
41 multivariate analysis, regimen including platinum plus PEM, a performance status > 0,  
42 non-epithelial histological type, and stage IV disease predicted poor OS in CTx patients.  
43 As hazard ratios of individual risk factors were approximately similar, a prognostic  
44 index for OS was constructed by counting the risk factors. Median OS in CTx patients  
45 decreased by each 1-point increase in this count: 1030 days for zero; 658 days for one;

373 days for two; 327 days for three; 125 days for four. Internal validation using the bootstrapping technique showed robustness of the model (c-index, 0.677; 95% Confidence Interval [CI], 0.624-0.729). Further, the discrimination was consistent in BSC patients (c-index, 0.799; 95% CI, 0.725-0.874).

## Conclusions

This novel index can provide clinicians and MPM patients with a better framework for discussing prognosis at the time of diagnosis.

## A mini-abstract:

We developed a new prognostic index for malignant pleural mesothelioma. The index reflects the recent real-world data. The index showed better discrimination than previous index.

## Keywords:

*Malignant pleural mesothelioma – pemetrexed - best supportive care - prognostic index - palliative care.*

## 63 INTRODUCTION

64 Malignant pleural mesothelioma (MPM) used to be a rare malignancy of the  
65 mesothelium. In recent years, the incidence of this disease has increased, and this trend  
66 will likely continue worldwide over the next decade (1).

67 Despite recent advancements in treatment, surgery, radiotherapy and chemotherapy or  
68 multimodality therapy has not be proven to be curative (2–4) . For the majority of  
69 patients, treatment options are limited to palliative chemotherapy and best supportive  
70 care (BSC) (5).

71 In oncologic palliative care, early determinations of prognosis play an important role in  
72 guiding end-of-life care and efforts designed to improve patients' quality of life (6, 7).

73 To determine the prognosis of patients with MPM, four prognostic indices (PI) have  
74 been developed; one by the Cancer and Leukaemia Group B (CALGB) (8), and three by  
75 the European Organization for Research and Treatment of Cancer (EORTC) (9–11).

76 While the first two PIs from EORTC can indicate either a favorable or an unfavorable  
77 outcome, neither can predict the duration of survival, which means both are impractical  
78 when discussing life expectancy with a patient. The CALGB PI is complex to use,  
79 because it has various cutoffs to consider. Above all, these PIs are based on clinical trial  
80 data and may not be applicable to the clinical setting. Further, they do not incorporate

information regarding pemetrexed, which can improve overall survival (OS), and does not incorporate recent advancements in supportive care (3, 12–14). Therefore, while existing PIs might be useful for researchers in deciding which patients to include in clinical trials, these systems are less useful for clinicians who need to discuss prognoses with their MPM patients.

The purpose of this study was to provide a new PI for OS in MPM patients who underwent treatment with pemetrexed or best supportive care in a recent clinical setting.

## **Materials and methods**

### **Study design and patients**

A retrospective cohort study was performed, covering the period between April 1<sup>st</sup>, 2007 and March 31<sup>st</sup>, 2013. The cohort was defined as all patients with histologically proven (15) MPM at either one of two tertiary hospitals that serve the South Hanshin medical region, which is an area of high MPM incidence area in Japan (16).

Patients who had more than one cancer, underwent autopsy, or who received palliative chemotherapy without pemetrexed were excluded, Because our purpose is to develop a new PI in MPM patients who underwent treatment with pemetrexed which is the “standard of care” (5). Patients who had received chemotherapy or radiotherapy before

diagnosis, trimodal therapy, or surgical therapy extra-pleural pneumonectomy or  
pleurectomy or decortication) were excluded to avoid confounding influences (17).

## **Definitions of prognostic variables**

Potential prognostic factors that were analyzed included: histological subtypes (15),  
International Mesothelioma Interest Group stage (18), chemotherapy regimen, age,  
gender, Eastern Cooperative Oncology Group performance status (PS) (19), subjective  
symptoms, smoking history, asbestos exposure history, comorbidities (Charlson score  
(20)) and baseline blood or effusion parameters at the time of diagnosis.

## **Primary outcomes measurement**

The primary outcome endpoint was OS, as defined by the length of time from the date  
of diagnosis to death. Patients who had not died or who were lost to follow-up were  
censored when they were last known to be alive before September 1<sup>st</sup>, 2013.

## **Statistical analyses**



115 We developed the PI in those who were treated with chemotherapy to minimize the bias  
116 due to confounding by indication (21). We also evaluated the applicability of the PI in  
117 those that received BSC.

118 In derivation, step continuous and nominal prognostic variables were dichotomized  
119 according to previous studies (8, 9, 11, 21–27). OS was estimated using the  
120 Kaplan–Meier method. The log-rank tests for each prognostic factor were used for  
121 univariate analyses. The Cox proportional hazards model was used for multivariate  
122 analyses. The Akaike’s information criterion (AIC), Schwartz’s Bayesian information  
123 criterion (BIC), and Harrell’s c index (c-index) were used for the discrimination of the  
124 model. A final model was chosen based on both clinical and statistical significance. We  
125 compared the discrimination of our index with the EORTC prognostic index (9) and the  
126 progression-free index of EORTC (11).

127 Calibration curves showing agreement between observed and predicted outcomes over a  
128 range of predicted probabilities were drawn. We also drew Cox-Snell residuals and  
129 measured Moreau, O’Quigley, and Lellouch statistics (28). We drew log-log hazards  
130 curves and tested the proportional hazard assumption. The bootstrapping technique was  
131 used for the internal validation (for 500 replications (29)).

132 We carried out sensitivity analysis using multiple imputation for variants with clinically  
133 significance. Two-sided p values < 0.05 were considered to indicate statistical  
134 significance. We used Stata® ver. 13.0 (Stata Corp., College Station, TX).

135

## 136 **Ethical considerations**

137 This study was performed according to the Declaration of Helsinki and the Ethical  
138 Guidelines for Epidemiological Research by the Japanese Ministry of Health, Labour  
139 and Welfare. The protocol for the study was approved by the Ethics Committee of  
140 Kyoto University Graduate School and Faculty of Medicine (E1883). The protocol was  
141 registered in the University Hospital Medical Information Network Clinical Trials  
142 Registry with the number: UMIN000011733.

143

## 144 **Results**

145 This study included 228 patients who were treated with chemotherapy with pemetrexed  
146 and 55 patients who received BSC (Figure 1). Patient characteristics are shown in Table  
147 1. Survival curves for each group are shown in the Figure 2.

148 The median lengths of follow-up were 345.5 days for the chemotherapy group and 250  
149 days for the BSC group. During the follow-up period, 161 patients (70.6%) died in  
150 chemotherapy group, and 40 patients (72.7%) died in the BSC group, respectively.

151 Univariate survival analyses are also shown in Table 1. Fifteen parameters were  
152 significantly correlated with OS according to univariate analyses: asbestos exposure, PS,  
153 dyspnea, anorexia, chest pain, body weight (BW) loss, fever, histological type, Stage,  
154 Regimen, white blood cell (WBC), platelet (Plt) count, C-reactive protein (CRP),  
155 Lactate dehydrogenase (LDH), and cytokeratin-19 fragment (CYFRA).

156 Because of the theoretical collinearity of symptom variables, we chose only PS with  
157 respect to clinical relevance. We repeated the multivariate analysis while analyzing  
158 WBC, Plt, and CRP, separately, because of the collinearity of inflammatory variables.

159 The discrimination for PS, Asbestos Exposure, Histology, Stage, Regimen, LDH, and  
160 CYFRA were 823 (AIC), 844 (BIC), and 0.714 (c-index). The discrimination for seven  
161 variables with WBC were 821 (AIC), 845 (BIC), and 0.726 (c-index). The  
162 discrimination for six variables with CRP were 825 (AIC), 849 (BIC), and 0.715  
163 (c-index). The discrimination for six variables with Plt were 824 (AIC), 848 (BIC), and  
164 0.711 (c-index). We entered WBC into a stepwise backward Cox proportional hazards  
165 model (Table 2). PS, histology, stage, and regimen remained significant after the

166 multivariate analysis. Hazard ratios of individual risk factors were 1.82-2.25. Therefore,  
167 a PI for the OS was constructed using a simple count of the number of risk factors  
168 (Table 3). The median OS of each category is shown in Table 4.

169 We calculated the discrimination of the rPHS (regimen, PS, Histology, or Stage) index.  
170 The c-index was 0.677. After 500 bootstrap replications from the original patients, the  
171 95% confidence interval (CI) of the c-index of the PHS score was 0.624-0.729.

172 We calculated the c-index for the EORTC prognostic index (9), which was 0.569. The  
173 difference between the two indices persisted after bootstrap replications (0.108; 95%CI,  
174 0.053-0.163). We also calculated the c-index for the progression-free index of the  
175 EORTC (11), which was 0.552. The difference between the two indices persisted after  
176 bootstrap replications (0.125, 95%CI, 0.082-0.166).

177 There was good calibration of the model, with close agreement between observed and  
178 predicted OS (Figure S1), and also with close agreement between Cox-Snell residuals  
179 and the 45-degree slope (Figure S2). The Moreau, O'Quigley, and Lellouch test showed  
180 that the model fit of the Cox regression model was adequate ( $p = 0.38$ ).

181 We drew log-log hazards curves for the CTx group which were parallel (Figure S3). The  
182 p value of the test for the proportional hazard assumption was 0.07.

183 We carried out sensitivity analysis using multiple imputation to create and analyze 10  
184 multiply imputed datasets. We imputed only PS with regards to clinical significance.  
185 These estimates and their standard errors were combined using Rubin's rules (30). The  
186 results showed consistency (Table 4). The discrimination was also consistent in the BSC  
187 group (c-index, 0.799; 95%CI, 0.725-0.874).

188

## 189 Discussion

190 We developed a new PI for patients with MPM that predicts median OS, incorporates  
191 pemetrexed information, and incorporates recent advancements in supportive care in the  
192 normal clinical setting. The rPHS index is obtained by a simple count of the risk factors  
193 (regimen including platinum plus PEM, PS>0, non-epithelial histology, and stage>3).  
194 The index can stratify patients into four different prognostic groups with different  
195 median survivals. The index has good discrimination for those treated with pemetrexed  
196 group as well as those treated with BSC.

197 Patients with advanced cancer often want to know their prognosis (31). One study (32)  
198 reported that patients with advanced cancer have an overwhelming preference for an  
199 opportunity to prepare for the end of life. They want to know that their families are  
200 prepared for their death, which often includes having finances in order, and for patients,

201    having funeral arrangements planned. They want to have the opportunity to resolve  
202    unfinished business, remember personal accomplishments, and to say goodbye to  
203    important people. In order to allow these patients to direct their energies to these matters,  
204    it is important to provide them with accurate information regarding their prognosis. In  
205    fact, early palliative care, including early accurate perceptions of prognosis, has  
206    improved the quality of life and possibly the OS of patients with advanced cancer (6).  
207    We believe that the present findings will influence the usual care of MPM patients for  
208    several reasons. When one patient diagnosed with MPM and decided to treat with  
209    pemetrexed-regimen, the patient and their physician can discuss based on the  
210    median OS of the rPHS index. Without the index we discussed the prognosis based  
211    on the median survival time from the trial or the cohort study. Our PI consists of  
212    variables frequently used in usual care of MPM patients. Indeed, PS, histology, and  
213    stage are well-known prognostic factors in previous studies (8–11, 33) and are  
214    components of the evaluation at the time of initial diagnosis (34). Further, our PI can be  
215    calculated easily by simple counting; calculators are not necessary, and our PI has more  
216    discriminatory power than the EORTC PI (9), which is one of the best-known clinical  
217    PIs. We note that the distribution of median age and OS were different when comparing  
218    previous reports (8–11) and our CTx cohort; our study included more elderly patients

219 (67.7 versus 58-62 years), and our study included patients with relatively better  
220 prognoses (11.5 versus 5-12.6 months). The cohort of our study ensures the  
221 generalizability of the findings, because the two hospitals cover the South Hanshin  
222 medical region and any patients with MPM in this region will visit one of these two  
223 hospitals. So, the participants in the present study are a good representation of patients  
224 with MPM. We included only patients with histologically proven MPM and not those  
225 with only cytologically proven MPM. Because there is morphologic overlap between  
226 benign reactive mesothelial cells and malignant cells of mesothelioma (15), it is not  
227 recommended to make a diagnosis of mesothelioma based on cytology alone (34). We  
228 think that this restriction ensured our study result.

229 Our cohort consisted of patients treated with BSC. For the small number of BSC  
230 participants we didn't develop another index for BSC patients, but validated PHS index.  
231 The discrimination was good (c-index, 0.799; 95% CI, 0.725-0.874). No previous study  
232 has validated a PI in patients treated with BSC. This information will be useful for  
233 discussions regarding prognosis between clinicians and their patients.

234 Since 1998, several PIs have been described. In contrast to our PHS index, other PIs  
235 were based on clinical trial data. Therefore, in the context of usual care, our PHS index  
236 might be more widely applicable than other PIs. We cannot compare our PI with the PI

237 of Bottomley (23) because we did not evaluate patients with the EORTC LC13 or  
238 QLQ-C30. Their PI's c-index was 0.66. The point estimation was similar to that of our  
239 PI. Pass (27) reported stage, histology, sex, age, treatment, adjuvant treatment, platelets  
240 and WBC are clinical prognostic factor except for PS. We think this discrepancy may  
241 reflect the difference of target population. We excluded those received surgery, but  
242 Pass's target population is those received either palliative or potentially curative surgery.  
243 There are several limitations in the study. First, this was a retrospective study with a  
244 substantial number of missing PS data, so we performed sensitivity analysis using  
245 multiple imputation. The result confirms the robustness of our model. Second, we were  
246 not able to know the reason why each patient treated with the modality because this is a  
247 retrospective study and treatment allocations were not protocol based. To clarify the  
248 preferences for treatment in MPM patients prospective qualitative and quantitative  
249 studies will be needed (36). But this limitation reflects the normal clinical setting. Third,  
250 we assessed internal validation with the bootstrap method, but the sample size of this  
251 study did not allow for external validation, so validation studies are needed.

252 We developed a new PI using PS, histology, and stage for MPM patients treated with  
253 chemotherapy or BSC. This PI will allow better discussion between clinicians and



254 patients with regards to prognosis. Further prospective studies using this PI are

255 warranted.

256

## 257 **Acknowledgements**

258 This study was partially supported by the Pfizer Health Research Foundation without

259 restriction of publication. We thank the following individuals for collecting data: Naoya

260 Ito, Makio Kondo, and Nobuko Maehashi.

261

## 262 **Conflict of interest statement**

263 None declared.

## 264 **References**

265 1. Robinson BM. Malignant pleural mesothelioma: an epidemiological perspective.  
266 Ann. Cardiothorac. Surg. 2012; 1(4):491–6.

267 2. Cao CQ, Yan TD, Bannon PG, McCaughan BC. A systematic review of  
268 extrapleural pneumonectomy for malignant pleural mesothelioma. J. Thorac.  
269 Oncol. 2010; 5(10):1692–703.

270 3. Van Schil PE, Opitz I, Weder W et al. Multimodal management of malignant  
271 pleural mesothelioma: where are we today? Eur. Respir. J. 2014; 44(3):754–64.

272 4. Bovolato P, Casadio C, Billè A et al. Does surgery improve survival of patients  
273 with malignant pleural mesothelioma?: a multicenter retrospective analysis of  
274 1365 consecutive patients. J. Thorac. Oncol. 2014; 9(3):390–6.

- 275 5. Boons CCLM, VAN Tulder MW, Burgers JA et al. The value of pemetrexed for  
276 the treatment of malignant pleural mesothelioma: a comprehensive review.  
277 Anticancer Res. 2013; 33(9):3553–61.
  
- 278 6. Temel JS, Greer JA, Muzikansky A et al. Early palliative care for patients with  
279 metastatic non-small-cell lung cancer. N. Engl. J. Med. 2010; 363(8):733–42.
  
- 280 7. Temel JS, Greer J a, Admane S et al. Longitudinal perceptions of prognosis and  
281 goals of therapy in patients with metastatic non-small-cell lung cancer: results of  
282 a randomized study of early palliative care. J. Clin. Oncol. 2011;  
283 29(17):2319–26.
  
- 284 8. Herndon JE, Green MR, Chahinian a P et al. Factors predictive of survival  
285 among 337 patients with mesothelioma treated between 1984 and 1994 by the  
286 Cancer and Leukemia Group B. Chest 1998; 113(3):723–31.
  
- 287 9. Curran D, Sahmoud T, Therasse P et al. Prognostic factors in patients with  
288 pleural mesothelioma: the European Organization for Research and Treatment of  
289 Cancer experience. J. Clin. Oncol. 1998; 16(1):145–52.
  
- 290 10. Maione P, Perrone F, Gallo C et al. Pretreatment quality of life and functional  
291 status assessment significantly predict survival of elderly patients with advanced  
292 non-small-cell lung cancer receiving chemotherapy: a prognostic analysis of the  
293 multicenter Italian lung cancer in the elderly s. J. Clin. Oncol. 2005;  
294 23(28):6865–72.
  
- 295 11. Francart J, Vaes E, Henrard S et al. A prognostic index for progression-free  
296 survival in malignant mesothelioma with application to the design of phase II  
297 trials: A combined analysis of 10 EORTC trials. Eur. J. Cancer 2009;  
298 45(13):2304–2311.
  
- 299 12. Vogelzang NJ, Rusthoven JJ, Symanowski J et al. Phase III study of pemetrexed  
300 in combination with cisplatin versus cisplatin alone in patients with malignant  
301 pleural mesothelioma. J. Clin. Oncol. 2003; 21(14):2636–44.
  
- 302 13. Musk a W, Olsen N, Alfonso H et al. Predicting survival in malignant  
303 mesothelioma. Eur. Respir. J. Off. J. Eur. Soc. Clin. Respir. Physiol. 2011;  
304 38(6):1420–4.

- 305 14. Gemba K, Fujimoto N, Aoe K et al. Treatment and survival analyses of  
306 malignant mesothelioma in Japan. *Acta Oncol. (Madr)*. 2013; 52(4):803–808.
  
- 307 15. Husain AN, Colby T, Ordonez N et al. Guidelines for pathologic diagnosis of  
308 malignant mesothelioma: 2012 update of the consensus statement from the  
309 International Mesothelioma Interest Group. *Arch. Pathol. Lab. Med.* 2013;  
310 137(5):647–67.
  
- 311 16. Kurumatani N, Kumagai S. Mapping the risk of mesothelioma due to  
312 neighborhood asbestos exposure. *Am. J. Respir. Crit. Care Med.* 2008;  
313 178(6):624–9.
  
- 314 17. Salas M, Hofman A, Stricker BH. Confounding by indication: an example of  
315 variation in the use of epidemiologic terminology. *Am. J. Epidemiol.* 1999;  
316 149(11):981–3.
  
- 317 18. Rusch VW. A proposed new international TNM staging system for malignant  
318 pleural mesothelioma. From the International Mesothelioma Interest Group.  
319 *Chest* 1995; 108(4):1122–8.
  
- 320 19. Oken MM, Creech RH, Tormey DC et al. Toxicity and response criteria of the  
321 Eastern Cooperative Oncology Group. *Am. J. Clin. Oncol.* 1982; 5(6):649–55.
  
- 322 20. Charlson ME, Pompei P, Ales KL, MacKenzie CR. A new method of classifying  
323 prognostic comorbidity in longitudinal studies: development and validation. *J.*  
324 *Chronic Dis.* 1987; 40(5):373–83.
  
- 325 21. Psaty BM, Siscovick DS. Minimizing bias due to confounding by indication in  
326 comparative effectiveness research: the importance of restriction. *JAMA* 2010;  
327 304:897–898.
  
- 328 22. Ryan CW, Herndon J, Vogelzang NJ. A review of chemotherapy trials for  
329 malignant mesothelioma. *Chest* 1998; 113(1 SUPPL.):66S–73S.
  
- 330 23. Metintas M, Metintas S, Ucgun I et al. Prognostic factors in diffuse malignant  
331 pleural mesothelioma: Effects of pretreatment clinical and laboratory  
332 characteristics. *Respir. Med.* 2001; 95(10):829–835.

- 333 24. Bottomley A, Coens C, Efficace F et al. Symptoms and patient-reported  
334 well-being: Do they predict survival in malignant pleural mesothelioma? A  
335 prognostic factor analysis of EORTC-NCIC 08983: Randomized phase III study  
336 of cisplatin with or without raltitrexed in patients with malignant pleural. *J. Clin.*  
337 *Oncol.* 2007; 25(36):5770–5776.
- 338 25. Tanrikulu AC, Abakay A, Kaplan MA et al. A clinical, radiographic and  
339 laboratory evaluation of prognostic factors in 363 patients with malignant pleural  
340 mesothelioma. *Respiration.* 2010; 80(6):480–7.
- 341 26. Kao SCH, Pavlakis N, Harvie R et al. High blood neutrophil-to-lymphocyte ratio  
342 is an indicator of poor prognosis in malignant mesothelioma patients undergoing  
343 systemic therapy. *Clin. Cancer Res.* 2010; 16(23):5805–5813.
- 344 27. Pinato DJ, Mauri FA, Ramakrishnan R et al. Inflammation-based prognostic  
345 indices in malignant pleural mesothelioma. *J. Thorac. Oncol.* 2012;  
346 7(3):587–594.
- 347 28. Pass HI, Giroux D, Kennedy C et al. Supplementary prognostic variables for  
348 pleural mesothelioma: a report from the IASLC staging committee. *J. Thorac.*  
349 *Oncol.* 2014; 9(6):856–64.
- 350 29. A Global Goodness-of-Fit Statistic for the Proportional Hazards Model. *J. R. Stat.*  
351 *Soc. Ser. C* 1985; 34(3):212–218.
- 352 30. Steyerberg EW, Bleeker SE, Moll H a et al. Internal and external validation of  
353 predictive models: A simulation study of bias and precision in small samples. *J.*  
354 *Clin. Epidemiol.* 2003; 56(5):441–447.
- 355 31. Rubin. *Multiple Imputation for Nonresponse in Surveys*, Hoboken, NJ, USA:  
356 John Wiley & Sons, Inc., 1987.
- 357 32. Degner LF, Kristjanson LJ, Bowman D et al. Information needs and decisional  
358 preferences in women with breast cancer. *JAMA* 1997; 277(18):1485–92.
- 359 33. Steinhauser KE, Christakis N a., Clipp EC et al. Preparing for the end of life:  
360 Preferences of patients, families, physicians, and other care providers. *J. Pain*  
361 *Symptom Manage.* 2001; 22(3):727–737.

- 362 34. Pass HI, Giroux D, Kennedy C et al. Supplementary Prognostic Variables for  
363 Pleural Mesothelioma: A Report from the IASLC Staging Committee. *J. Thorac.*  
364 *Oncol.* 2014; 9(6):856–64.
- 365 35. Scherpereel A, Astoul P, Baas P et al. Guidelines of the European Respiratory  
366 Society and the European Society of Thoracic Surgeons for the management of  
367 malignant pleural mesothelioma. *Eur. Respir. J.* 2010; 35(3):479–495.
- 368 36. Muhlbcher AC, Bethge S. Patients ’ preferences : a discrete-choice experiment  
369 for treatment of non-small-cell lung cancer. *Eur. J. Heal. Econ.* 2014; Aug  
370 19.:Epub ahead of print.
- 371
- 372

Table 1 Patient characteristics and results of univariate analyses of OS

Clinical factors	Chemotherapy (n = 228), N (%)	Median OS (days)	95% CI	p value	BSC (n=55), n (%)
<b>Age (years)</b> <b>mean±SD</b>	67.7±8.2				74.5±9.6
<b>Age (years)</b>					
75>	181 (79.4)	512	375-562	0.2000	24 (43.6)
75≤	47 (20.6)	366	190-441		31 (56.4)
<b>Gender</b>					
Female	39 (17.1)	514	314-699	0.4700	13 (23.6)
Male	189 (82.9)	432	359-524		42 (76.4)
<b>Smoke</b>					
Never	65 (30.0)	524	366-624	0.4100	19 (38.0)
Current / Ever	152 (70.0)	425	327-524		31 (62.0)
Missing	11				5
<b>Charlson comorbidity index</b>					
<2	205 (89.9)	461	372-533	0.5100	42 (76.4)
2≤	23 (10.1)	366	224-1213		13 (23.6)
<b>Asbestos exposure</b>					
No	28 (12.4)	710	327-1213	0.0480	
Yes	197 (87.6)	397	353-511		14 (26.4)
Missing	3				39 (73.6)
<b>PS</b>					2
0	37 (22.4)	926	524-1372	0.0014	
1≤	128 (77.6)	434	362-562		19 (41.3)
Missing	63				37 (58.7)
<b>Dyspnea</b>					9
No	52 (33.8)	658	524-1030	0.0003	13 (31.7)
Yes	102 (66.2)	425	319-512		28 (68.3)
Missing	74				14
<b>Anorexia</b>					
No	145 (82.4)	524	432-654	0.0001	29 (55.8)

Yes	31 (17.6)	296	166-373		23 (44.2)
Missing	52				3
<b>Chest pain</b>					
No	58 (39.2)	648	511-926	0.0007	16 (43.2)
Yes	90 (60.8)	353	263-432		21 (56.8)
Missing	80				18
<b>BW loss</b>					
No	96 (70.7)	566	512-804	0.0001	20 (60.6)
Yes	41 (30.0)	299	177-425		13 (39.4)
Missing	91				22
<b>Fever</b>					
No	92 (76.7)	524	397-648	0.0280	38 (92.7)
Yes	28 (23.3)	353	223-518		3 (7.3)
Missing	108				14
<b>Histological type</b>					
Epithelial	149 (65.4)	545	493-640	0.0000	17 (30.9)
Non-epithelial	79 (34.7)	277	221-330		38 (69.1)
<b>Stage</b>					
I-III	133 (58.3)	549	461-658	0.0000	30 (54.5)
IV	95 (41.7)	327	242-375		25 (45.5)
<b>Regimen</b>					
Platinum plus PEM	205 (89.9)	221	373-547	0.0007	
PEM monotherapy	23 (10.1)	499	86-425		
<b>WBC (/μl)</b>					
8300>	160 (70.5)	512	391-598	0.0400	36 (65.5)
8300≤	67 (29.5)	359	238-501		19 (34.5)
Missing	1				
<b>Neutro/lymph</b>					
5>	182 (82.4)	445	368-549	0.0600	35 (64.8)
5≤	39 (17.7)	362	188-514		19 (35.2)
Missing	7				1
<b>Hb (g/dL)</b>					
10≤	218 (96)	445	372-544	0.0700	45 (81.8)
10>	9 (4.0)	224	66-526		10 (18.2)
Missing	1				

<b>Plt (10<sup>5</sup>/μl)</b>						
40>	188 (82.8)	461	373-549	0.0100	42 (76.4)	
40≤	39 (17.2)	327	176-526		13 (23.6)	
Missing	1					
<b>ALP (IU/l)</b>						
Abnormal	32 (14.9)	397	228-562	0.93	0 (0.0)	
Normal	183 (85.1)	441	362-544		53 (100.0)	
Missing	13				2	
<b>LDH (IU/L)</b>						
Abnormal	26 (11.6)	493	375-544	0.011	11 (20.4)	
Normal	198 (88.4)	242	87-603		43 (79.6)	
Missing	4				1	
<b>CRP (mg/dl)</b>						
5>	189 (83.3)	461	373-549	0.0076	40 (72.7)	
5≤	38 (16.7)	359	167-518		15 (27.3)	
Missing	1					
<b>CEA (ng/ml)</b>						
5>	200 (94.3)	338	156-NE	0.7800	45 (95.7)	
5≤	12 (5.7)	441	366-526		2 (4.3)	
Missing	16				8	
<b>CYFRA (ng/ml)</b>						
3.5>	162 (75)	512	375-598	0.0090	22 (48.9)	
3.5≤	54 (25)	368	242-445		23 (51.1)	
Missing	12				10	
<b>Pleural glucose (mg/dl)</b>						
40>	21(22.3)	511	156-710	0.2200	10 (30.3)	
40≤	73(77.7)	373	319-547		23 (69.7)	
Missing	134				22	

Abbreviations: N, number; OS, overall survival; SD, standard deviation; CI, confidence interval; BSC, best supportive care; PS, Eastern Cooperative Oncology Group performance status; BW, body weight; PEM, pemetrexed; WBC, white blood cell; Neutro, neutrocyte; Lymph, lymphocyte; Hb, hemoglobin; Plt, platelet; ALP, alkaly phosphatase; LDH, lactate dehydrogenase; CRP, C-reactive protein; CEA, carcinoembryonic antigen; CYFRA, cytokeratin-19 fragment.



Table 2 Backward Cox proportional hazards model

Clinical factors	HR	95%CI
<b>PS</b>		
0	1	
1≤	2.40	1.36-4.23
<b>Asbestos exposure</b>		
no	1	
yes	1.64	0.75-3.58
<b>Histological type</b>		
Epithelial	1	
Non-epithelial	2.16	1.40-3.32
<b>Regimen</b>		
Platinum doublet	1	
Pemetrexed only	3.18	1.59-6.39
<b>Stage</b>		
I-III	1	
IV	1.57	1.03-2.39
<b>LDH</b>		
Normal	1	
Abnormal	1.46	0.71-2.99
<b>CYFRA (ng/ml)</b>		
3.5>	1	
3.5≤	1.10	0.69-1.76
<b>WBC (/μl)</b>		
8300>	1	
8300≤	1.56	0.99-2.45

Abbreviations: HR, hazard ratio; CI, confidence interval; PS, Eastern Cooperative Oncology Group performance status; LDH, lactate dehydrogenase; CYFRA, cytokeratin-19 fragment; WBC, white blood cell.

Table 3 Final model

Clinical factors	HR	95% CI	Score
<b>PS</b>			
0	1		1
1≤	2.06	1.22-3.44	
<b>Histological type</b>			
Epithelial	1		1
Non-epithelial	2.15	1.41-3.26	
<b>Stage</b>			
I-III	1		1
IV	1.82	1.23-2.69	
<b>Regimen</b>			
Platinum plus PEM	1		1
PEM monotherapy	2.25	1.16-4.36	

Abbreviations: HR, hazard ratio; CI, confidence interval; PS, Eastern

Cooperative Oncology Group performance status; PEM, pemetrexed.

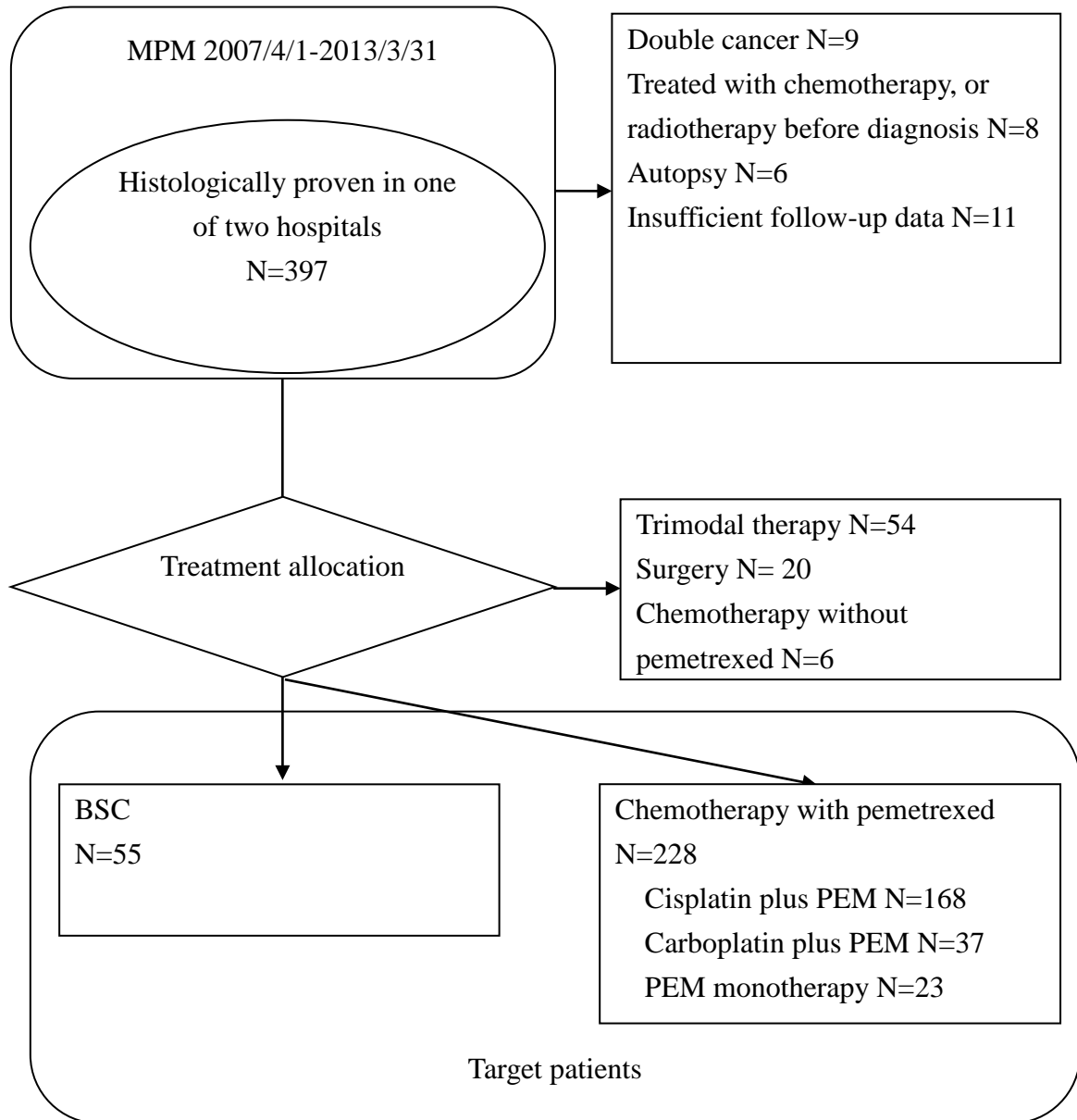
Table 4 The rPHS index for overall survival (sensitivity analysis)

Chemotherapy				Best supportive care		
Score	N	Median OS (days)	95% CI	N	Median OS (days)	95%CI
0	24 (28)	1030 (926)	661-1399(598-1253)			
1	57 (76)	658 (603)	444-872 (458-678)	6 (7)	573 (573)	530-616 (477-669)
2	56 (79)	373 (367)	223-522 (305-429)	15 (20)	408 (402)	178-638 (221-583)
3	22 (39)	327 (240)	189-465 (133-347)	11 (20)	250 (94)	11-489 (0-228)
4	5(6)	125(48)	16-234(0-184)	6 (8)	26 (34)	0-103 (0-126)

rPHS index = (if platinum + PEM 0, otherwise 1) + (if PS 0<, otherwise 0) + (if Histology non-epithelial, otherwise 0) + (if Stage=4, otherwise 0)

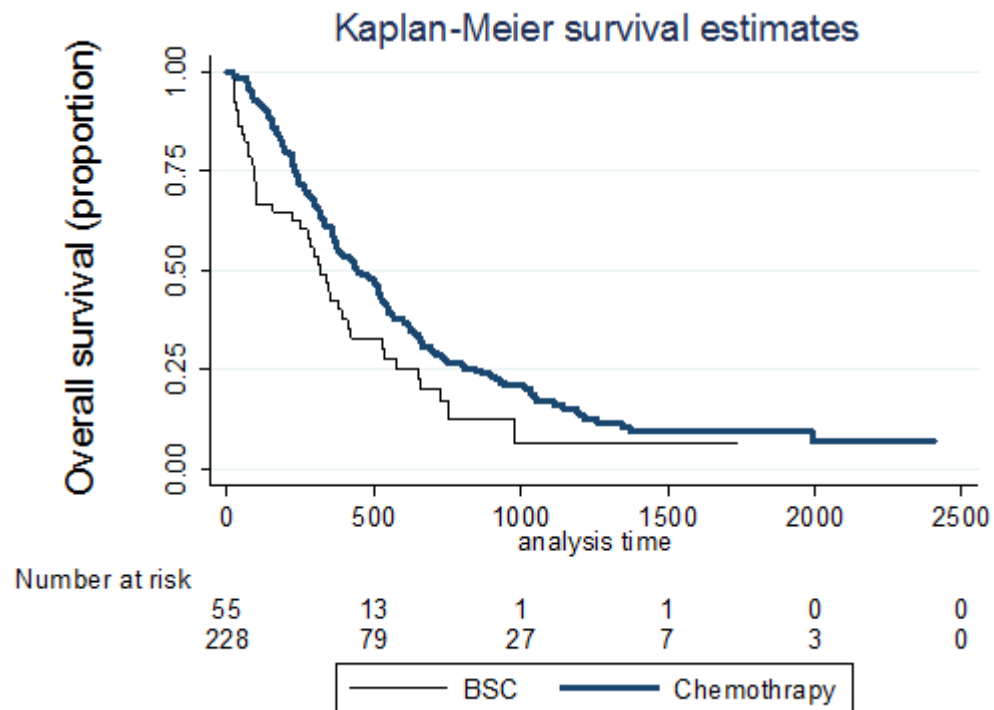
Abbreviations: N, number; OS, overall survival; CI, confidence interval; NE, not estimable; PS, Eastern Cooperative Oncology Group performance status; PEM, pemetrexed.

Figure 1 Patient flow diagram



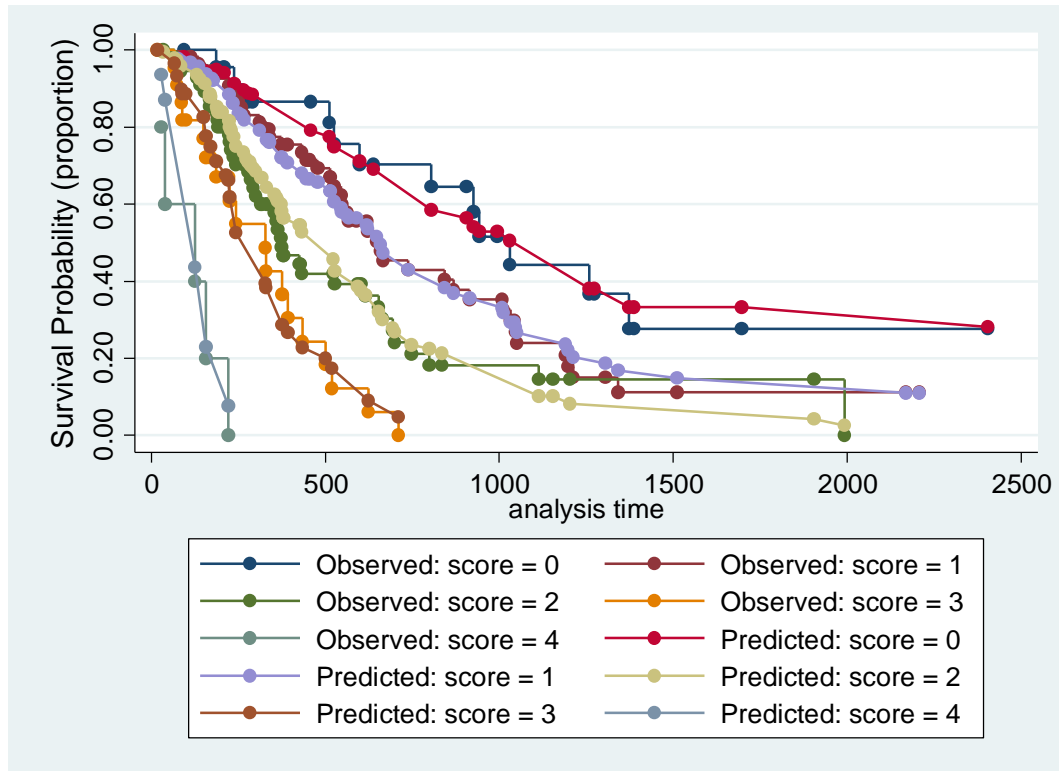
Abbreviations: N, number; MPM, malignant pleural mesothelioma; BSC, best supportive care.

Figure 2 Survival curve (days)



Abbreviations: BSC, best supportive care;

Figure S1 Calibration Kaplan-Meire curve of the rPHS index for chemotherapy group



Abbreviations: BSC, best supportive care;

Figure S2 Cox-Snell Residuals Graph

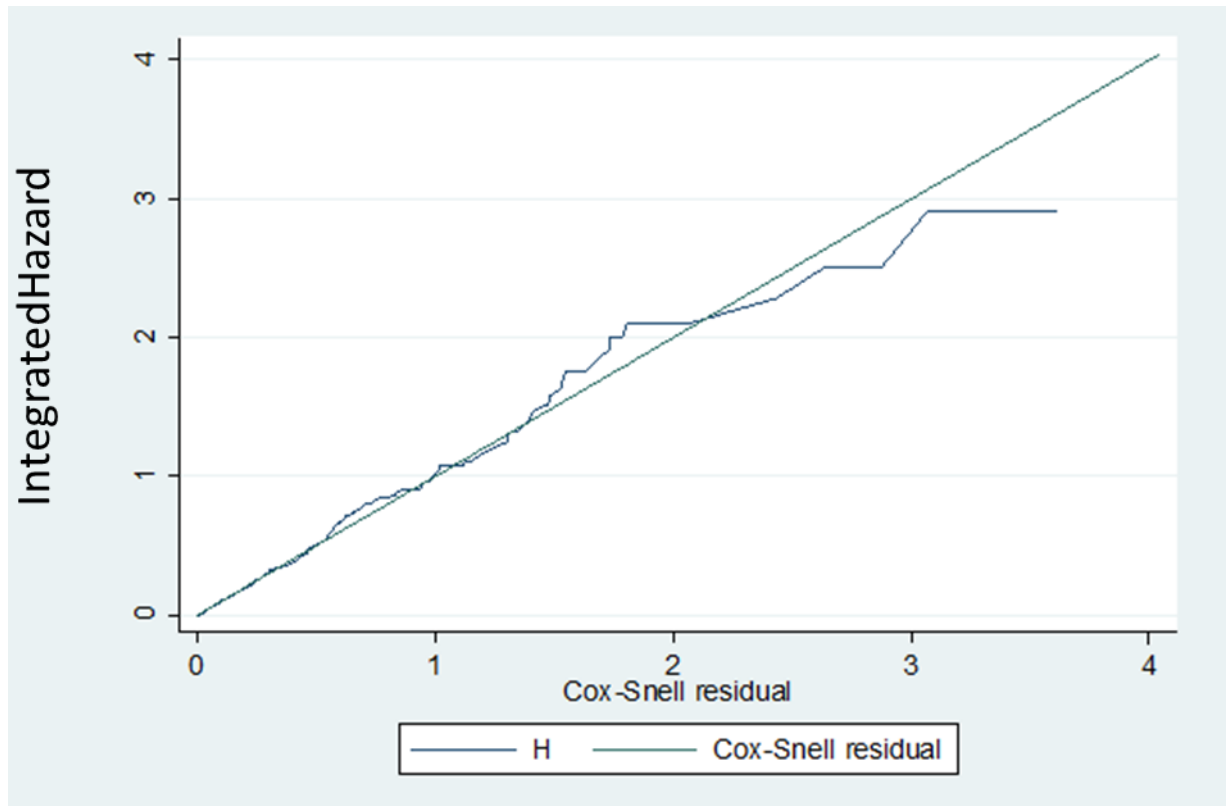


Figure S3 Cumulative hazards curves for the pemetrexed (CTx) group

